## EE5847: Information Theory

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Handout 4: Properties of the information measures - 1

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# 4.1 Recap

- The minimum rate for fixed-length compression of a memoryless source with distribution  $p_X$  is equal to the entropy H(X).
- The maximum rate of reliable communication over a DMC  $p_{Y|X}$  is equal to the capacity  $C = \max_{p_X} I(X;Y)$ .
- The optimal error exponent for classifying between two sources  $p_s, p_g$  is equal to the KL divergence  $D(p_s || p_q)$ .

The material covered in this handout is contained in Chapter 2 of the book by Cover and Thomas.

## 4.2 Relook of the information measures

Verify that we can write

$$H(X) = \mathbb{E}_X \log_2 \frac{1}{p_X(X)}$$

$$H(X,Y) = \mathbb{E}_{XY} \log_2 \frac{1}{p_{XY}(X,Y)}$$

$$H(X|Y) = \mathbb{E}_{XY} \log_2 \frac{1}{p_{X|Y}(X|Y)}$$

$$I(X;Y) = \mathbb{E}_{XY} \log_2 \frac{p_{XY}(X,Y)}{p_X(X)p_Y(Y)}.$$

We have already seen in previous classes that

$$I(X;Y) = H(X) - H(X|Y).$$

# 4.3 Properties

## 4.3.1 Properties

**Lemma 4.1.** *Entropy is nonnegative.* 

This follows from the fact that  $-x \log x > 0$  for all  $x \in (0,1)$ .

## 4.3.2 Chain rules

Consider the definition of joint entropy of two correlated random variables X, Y.

$$H(X,Y) = \mathbb{E}_{X,Y} \log_2 \frac{1}{p_{XY}(X,Y)}$$

$$= \mathbb{E}_{X,Y} \log_2 \frac{1}{p_X(X)p_{Y|X}(Y|X)}$$

$$= \mathbb{E}_{X,Y} \log_2 \frac{1}{p_X(X)} + \mathbb{E}_{X,Y} \log_2 \frac{1}{p_{Y|X}(Y|X)}$$
(4.1)

This gives us the following chain rule of entropy:

$$H(X,Y) = H(X) + H(Y|X) = H(Y) + H(X|Y). \tag{4.2}$$

and

$$H(X,Y|Z) = H(X|Z) + H(Y|X,Z).$$
 (4.3)

This will be useful in proving a number of chain rules.

#### 4.3.2.1 More variables

$$H(X_1, \dots, X_n) = \sum_{i=1}^n H(X_i | X_1, \dots, X_{i-1}).$$
(4.4)

This can be proved by repeatedly using (4.2) and (4.3).

$$H(X_1, \dots, X_n) = H(X_1) + H(X_2, \dots, X_n | X_1)$$
(4.5)

$$= H(X_1) + H(X_2|X_1) + H(X_3, \dots, X_n|X_1, X_2)$$
(4.6)

$$\vdots (4.7)$$

$$= H(X_1) + H(X_2|X_1) + \ldots + H(X_n|X_1, \ldots X_{n-1}). \tag{4.8}$$

#### 4.3.2.2 Chain rule of mutual information

The conditional mutual information is defined as

$$I(X;Y|Z) = H(X|Z) - H(X|Y,Z).$$

We can prove the following chain rule

$$I(X_1, \dots, X_n; Y) = \sum_{i=1}^n I(X_i; Y | X_1, \dots, X_{i-1})$$
(4.9)

This follows by repeatedly using the chain rule of entropy for each of the conditional entropy terms.

$$I(X_1, X_2; Y) = H(X_1, X_2) - H(X_1, X_2|Y)$$

$$= H(X_1) + H(X_2|X_1) - H(X_1|Y) - H(X_2|X_1, Y)$$

$$= I(X_1; Y) + I(X_2; Y|X_1).$$

#### 4.3.2.3 Chain rule for KL divergence

We can prove something similar here as well: For any pair of joint distributions  $p_{XY}$  and  $q_{XY}$  such that  $p_{XY}(x,y) = 0$  whenever  $q_{XY}(x,y) = 0$ , we have

$$D(p_{XY}||q_{XY}) = D(p_X||q_X) + D(p_{Y|X}||q_{Y|X}),$$

where we define

$$D(p_{Y|X}||q_{Y|X}) \stackrel{\text{def}}{=} \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_{Y|X}(x,y)}{q_{Y|X}(x,y)}.$$

This follows from definition.

$$D(p_{XY}||q_{XY}) = \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_{XY}(x,y)}{q_{XY}(x,y)}$$

$$= \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_X(x)p_{Y|X}(y|x)}{q_X(x)q_{Y|X}(y|x)}$$

$$= \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_X(x)}{q_X(x)} + \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_{Y|X}(y|x)}{q_{Y|X}(y|x)}$$

$$= \sum_{x} p_X(x) \log_2 \frac{p_X(x)}{q_X(x)} + \sum_{x,y} p_{XY}(x,y) \log_2 \frac{p_{Y|X}(y|x)}{q_{Y|X}(y|x)}$$

$$= D(p_X||q_X) + D(p_{Y|X}||q_{Y|X}).$$

## 4.4 Convex sets and functions

A set  $S \subset \mathbb{R}^m$  is said to be convex if the line segment joining any two points within S also lies in S. Formally, S is convex if

$$\alpha x_1 + (1 - \alpha)x_2 \in \mathcal{S}$$
 for all  $x_1, x_2 \in \mathcal{S}$  and  $\alpha \in [0, 1]$ .

Note that every closed interval  $[a, b] \subset \mathbb{R}$  is convex.

A function defined on a convex set S,  $f: S \to \mathbb{R}$  is convex if for all  $x, y \in S$  and  $\alpha \in [0, 1]$ , we have

$$f(\alpha x + (1 - \alpha)y) \le \alpha f(x) + (1 - \alpha)f(y).$$

The function is strictly convex if equality holds only for  $\alpha = 0, 1$ .

A function f is concave if -f is convex.

**Lemma 4.2.** A twice-differentiable function  $f : \mathbb{R} \to \mathbb{R}$  has nonnegative second derivative on [a,b] if and only if it is convex in [a,b]. If the second derivative is strictly positive in the interval, then it is also strictly convex and vice versa.

*Proof.* To show that the second derivative being nonnegative implies convexity, see Theorem 2.6.1 in Cover and Thomas.

For the other way around, recall from your calculus course that

$$f''(x) = \lim_{t \to 0} \frac{f(x+t) + f(x-t) - 2f(x)}{t^2}.$$

It therefore suffices to show that  $\frac{f(x+t)+f(x-t)-2f(x)}{t^2} \geqslant 0$  for all t > 0.

Now since f is convex,

$$f(x) = f\left(\frac{x+t}{2} + \frac{x-t}{2}\right) \le \frac{1}{2}f(x+t) + \frac{1}{2}f(x-t).$$

Rearranging, we get that

$$f(x+t) + f(x-t) - 2f(x) \ge 0.$$

This completes the proof. If the function is strictly convex, then the inequality is strict, and therefore the second derivative is positive.  $\Box$ 

A very useful inequality for convex functions is the following:

**Lemma 4.3** (Jensen's inequality). For any convex function  $f: \mathbb{R} \to \mathbb{R}$  and random variable X, we have

$$\mathbb{E}f(x) \geqslant f(\mathbb{E}X).$$

If f is strictly convex, then equality implies that X is a constant.

*Proof.* The Cover and Thomas book gives a proof specific to discrete random variables using mathematical induction. Here is a shorter proof:

Let  $c + \alpha x$  denote the tangent to f(x) at the point  $\mathbb{E}X$ . I claim that the following is true (Why?):

$$c + \alpha x \leq f(x)$$
.

Using this,

$$\mathbb{E}f(X) \geqslant \mathbb{E}(a+bX) = a+b\mathbb{E}X \tag{4.10}$$

However, the line intersects f(x) for  $x = \mathbb{E}X$ , and hence  $a + \mathbb{E}X = f(\mathbb{E}X)$ . This completes the proof.

Equality in (4.10) implies that f(x) = a + bx for all x having nonzero probability (or nonzero density for continuous rvs). If the random variable is not a constant, then it means that the second derivative of f is zero and hence it is not strictly convex.

Jensen's inequality is the basis for many other results in information theory and statistics.

#### 4.4.0.1 Nonnegativity of KL divergence

**Lemma 4.4.**  $D(p||q) \ge 0$  with equality iff p(x) = q(x) for all x such that p(x) > 0.

*Proof.* Let S denote the set of all x such that p(x) > 0. This is called the support of p. The main observation here is that  $\log_2(x)$  is a concave function of x, and we can use Jensen's inequality. To start with, consider

$$-D(p||q) = -\sum_{x \in S} p(x) \log_2 \frac{p(x)}{q(x)} = \sum_{x \in S} p(x) \log_2 \frac{q(x)}{p(x)}$$

Note that the last term can be written as  $\mathbb{E}_p \log_2 \frac{q(X)}{p(X)}$ , where  $X \sim p$ . Using Jensen's inequality,

$$-D(p||q) \le \log_2 \left( \sum_{x \in \mathcal{S}} p(x) \frac{q(x)}{p(x)} \right)$$

$$= \log_2 \left( \sum_{x \in \mathcal{S}} q(x) \right)$$
 
$$= \log_2(1) = 0.$$

Rearranging, we get  $D(p||q) \ge 0$ .

### 4.4.0.2 Implications of nonnegativity of KL divergence

 $I(X;Y) \geqslant 0.$ 

And equality holds in the above iff X, Y are independent. This follows from the fact that  $I(X; Y) = D(p_{XY} || p_X p_Y)$ .

 $D(p_{Y|X}\|q_{Y|X}) \geqslant 0$ 

with equality iff  $p_{Y|X}(y|x) = q_{Y|X}(y|x)$  for all (x,y) such that  $p_{Y|X}(y|x) > 0$ .

 $I(X;Y|Z) \geqslant 0$ 

• Conditioning reduces entropy

$$H(Y|X) \leq H(Y)$$

with equality iff X and Y are independent. This follows from I(X;Y) = H(Y) - H(Y|X).

 $H(X) \leq \log_2 |\mathcal{X}|.$ 

Consider  $q(x) = \frac{1}{|\mathcal{X}|}$  for all x and  $p(x) = p_X(x)$ . Simplifying D(p||q), we get

$$0 \leqslant D(p||q) = \log_2 |\mathcal{X}| - H(X).$$

 $H(X_1,\ldots,X_n)\leqslant \sum_{i=1}^n H(X_i).$ 

#### 4.4.0.3 Log-sum inequality

**Lemma 4.5.** Suppose that  $\alpha_1, \ldots, \alpha_k$  and  $\beta_1, \ldots, \beta_k$  are nonnegative numbers such that  $\alpha_i > 0$  whenever  $\beta_i > 0$ . Then,

$$\sum_{i} \alpha_{i} \log_{2} \frac{\alpha_{i}}{\beta_{i}} \geqslant \left(\sum_{i} \alpha_{i}\right) \log_{2} \frac{\left(\sum_{i} \alpha_{i}\right)}{\left(\sum_{i} \beta_{i}\right)}$$

and equality holds if and only if  $\alpha_i = \beta_i$  for all i such that  $\alpha_i > 0$ .

*Proof.* We will again use Jensen's inequality on the strictly convex function  $x \log x$ . Start with the LHS.

$$\sum_{i=1}^{k} \alpha_{i=1}^{k} \log_2 \frac{\alpha_i}{\beta_i} = \sum_{i=1}^{k} \beta_i \frac{\alpha_i}{\beta_i} \log_2 \frac{\alpha_i}{\beta_i}$$

$$= \left(\sum_{j=1}^{k} \beta_{j}\right) \left(\sum_{i=1}^{k} \frac{\beta_{i}}{\sum_{j=1}^{k} \beta_{j}} \frac{\alpha_{i}}{\beta_{i}} \log_{2} \frac{\alpha_{i}}{\beta_{i}}\right)$$

If we set  $p(i) = \beta_i / \sum_{j=1}^k \beta_j$ , and  $t_i = \alpha_i / \beta_i$ , then the above becomes

$$\sum_{i=1}^{k} \alpha_{i=1}^{k} \log_{2} \frac{\alpha_{i}}{\beta_{i}} = \left(\sum_{j=1}^{k} \beta_{j}\right) \sum_{i=1}^{k} p(i)(t_{i} \log_{2} t_{i})$$

$$\geqslant \left(\sum_{j=1}^{k} \beta_{j}\right) \left(\sum_{i=1}^{k} p(i)t_{i}\right) \log_{2} \left(\sum_{i=1}^{k} p(i)t_{i}\right)$$

$$\geqslant \left(\sum_{j=1}^{k} \beta_{j}\right) \left(\sum_{i=1}^{k} \frac{\beta_{i}}{\sum_{j=1}^{k} \beta_{j}} \frac{\alpha_{i}}{\beta_{i}}\right) \log_{2} \left(\sum_{i=1}^{k} \frac{\beta_{i}}{\sum_{j=1}^{k} \beta_{j}} \frac{\alpha_{i}}{\beta_{i}}\right)$$

$$= \left(\sum_{i} \alpha_{i}\right) \log_{2} \frac{\left(\sum_{i} \alpha_{i}\right)}{\left(\sum_{j} \beta_{j}\right)}$$

This can be generalized easily to the continuous case by replacing summations with integrals in the proofs (assuming that they all exist).

**Lemma 4.6.** Let f, g be nonnegative integrable functions on  $\mathbb{R}$  with  $\int_{x=-\infty}^{\infty} f(x)dx > 0$  and  $\int_{x=-\infty}^{\infty} g(x)dx > 0$ . Additionally, f(x) = 0 whenever g(x) = 0. Then,

$$\int_{x=-\infty}^{\infty} f(x) \log_2 \frac{f(x)}{g(x)} dx \geqslant \left( \int_{x=-\infty}^{\infty} f(x) dx \right) \log_2 \frac{\left( \int_{x=-\infty}^{\infty} f(x) dx \right)}{\left( \int_{x=-\infty}^{\infty} g(x) dx \right)}.$$