Averaging $Q(\|\mathbf{X}\|)$ for a Complex Gaussian Random Vector \mathbf{X} : A Novel Approach

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Abstract—In this paper, we compute $E[Q(\|\mathbf{X}\|]]$, where \mathbf{X} is an $n \times 1$ complex circularly Gaussian vector, $\|\mathbf{X}\|$ is the L^2 norm of \mathbf{X} and $E[\]$ is the expectation operator. This is done by finding the characteristic function of the decision variable and subsequently applying the inversion formula to obtain a one dimensional real integral expression. This integral is then converted to a contour integral which is evaluated using a variant of the Cauchy's integral formula to obtain an expression for $E[Q(\|\mathbf{X}\|]]$. We then provide some applications of the above result by obtaining expressions for error probabilities in fading channels.

I. INTRODUCTION

The average of a Q-function expression is of interest in finding general expressions for the probability of symbol/bit error in slowly fading communication channels, where the argument of the Q-function is a function of a random variable with a well defined probability density function that depends on the kind of fading experienced by the channel [1]. Another application is in finding the bit error probability of the desired user in a fading channel when multiuser detection is being employed at the receiver [2].

The average of a Q-function whose argument is proportional to the square root of a non-central chi-squared random variable with 2n degrees of freedom is obtained by deriving a recursion relation [3]. However, this approach involves integrating the product of the Q-function and the chi-square probability density function. [2]. For the central chi-square distribution, an approach using Craig's formula [4] and the moment generating function (MGF) of the random variable is outlined in [1]. This method is quite complicated but gives the most general closed form expression when the Q-function argument is proportional to the square root of a Nakagami-m [5] random variable.

In this paper, we propose a simple method to average the Q-function whose argument is the L^2 norm of a zero-mean complex circularly Gaussian vector using the characteristic function (CF) of a non-central chi-square distribution [5] and the inversion formula for the cumulative distribution function [6].

II. PROBLEM STATEMENT

Let X be an $n \times 1$ vector whose entries are complex circularly Gaussian random variables such that

$$E[\mathbf{X}] = \mathbf{m}, \qquad (1)$$

$$E[(\mathbf{X} - \mathbf{m})(\mathbf{X} - \mathbf{m})^{\dagger}] = 2\sigma^{2}\mathbf{I}_{n},$$

where $\{\dagger\}$ represents the complex conjugate-transpose operation and \mathbf{I}_n is the $n \times n$ identity matrix. Let $R = \|\mathbf{X}\|^2$, where $\|\mathbf{X}\|$ is the L^2 norm of the vector \mathbf{X} . Then R is non-central chi-square distributed with 2n degrees of freedom [5] and

$$E[Q(\|\mathbf{X}\|)] = E[Q(\sqrt{R})]. \tag{2}$$

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Let V be a Gaussian random variable with zero mean and unit variance. From [7], we obtain

$$\begin{split} E[Q(\sqrt{R})] &= E[P(V > \sqrt{R})] \\ &= \frac{1}{2}P(R - V^2 < 0) \\ &= \frac{1}{2}P(\Delta < 0), \end{split} \tag{3}$$

where $\Delta=R-V^2$. The characteristic function of Δ is given by

$$\Phi_{\Delta}(t) = E[e^{jt\Delta}] = E[e^{jtR}]E[e^{-jtV^{2}}]
= \Phi_{R}(t)\Phi_{V^{2}}(-t).$$
(4)

Since V is Gaussian, V^2 chi-square distributed. Hence, we obtain [5]

$$\Phi_R(t) = \frac{e^{\frac{j\|\mathbf{m}\|^2 t}{1 - 2j\sigma^2 t}}}{(1 - 2j\sigma^2 t)^n}, \qquad (5)$$

$$\Phi_{V^2}(-t) = \frac{1}{(1 + 2jt)^{\frac{1}{2}}}.$$

Substituting $\Phi_R(t)$ and $\Phi_{V^2}(-t)$ from (5) in (4),

$$\Phi_{\Delta}(t) = \frac{e^{\frac{j\|\mathbf{m}\|^2 t}{1 - 2j\sigma^2 t}}}{(1 - 2j\sigma^2 t)^n (1 + 2jt)^{\frac{1}{2}}}.$$
 (6)

Since

$$\frac{j\|\mathbf{m}\|^2 t}{1 - 2j\sigma^2 t} = -\frac{\|\mathbf{m}\|^2}{2\sigma^2} + \frac{\|\mathbf{m}\|^2}{2\sigma^2 (1 - 2j\sigma^2 t)},\tag{7}$$

(6) can be written as

$$\Phi_{\Delta}(t) = \frac{e^{-\frac{\|\mathbf{m}\|^2}{2\sigma^2}} e^{\frac{\|\mathbf{m}\|^2}{2\sigma^2(1-2j\sigma^2t)}}}{(1-2j\sigma^2t)^n (1+2jt)^{\frac{1}{2}}}.$$
 (8)

According to the inversion formula of Gil-Pelaez [6], the cumulative distribution function

$$F_{\Delta}(x) = P(\Delta < x)$$

$$= \frac{1}{2} + \frac{1}{2\pi i} \int_{0}^{\infty} \frac{e^{jtx}\Phi_{\Delta}(-t) - e^{-jtx}\Phi_{\Delta}(t)}{t} dt.$$

$$(9)$$

From (9), we have

$$P(\Delta < 0) = \frac{1}{2} + \frac{1}{2\pi j} \int_0^\infty \frac{\Phi_\Delta(-t) - \Phi_\Delta(t)}{t} dt$$
 (10)
$$= \frac{1}{2} + \frac{1}{2\pi j} \int_{-\infty}^\infty \frac{\Phi_\Delta(-t)}{t} dt.$$

From (8) and (10),

$$P(\Delta < 0) = \frac{1}{2} + \frac{e^{-\frac{\|\mathbf{m}\|^2}{2\sigma^2}}}{2\pi j} \int_{-\infty}^{\infty} \frac{e^{\frac{\|\mathbf{m}\|^2}{2\sigma^2(1+2j\sigma^2t)}}}{t(1-2jt)^{\frac{1}{2}}(1+2j\sigma^2t)^n} dt.$$
(11)

Expanding the exponential in the numerator of the integrand in (11) as a power series and interchanging the order of integration and summation, we obtain

$$P(\Delta < 0) = \frac{1}{2} + \frac{e^{-\frac{\left\|\mathbf{m}\right\|^{2}}{2\sigma^{2}}}}{2\pi j} \sum_{p=0}^{\infty} \frac{1}{p!} \left(\frac{\left\|\mathbf{m}\right\|^{2}}{2\sigma^{2}}\right)^{p} \int_{-\infty}^{\infty} \frac{dt}{t(1-2jt)^{\frac{1}{2}}(1+2j\sigma^{2}t)^{n+p}}.$$

It is easy to verify that

$$\frac{1}{t(1+2j\sigma^2t)^n} = \frac{1}{t} - 2j\sigma^2 \sum_{k=1}^n \frac{1}{(1+2j\sigma^2t)^k}.$$
 (13)

If we let

$$I_n = \int_{-\infty}^{\infty} \frac{dt}{t(1 - 2it)^{\frac{1}{2}} (1 + 2i\sigma^2 t)^n},$$
 (14)

(12) can be written as

$$P(\Delta < 0) = \frac{1}{2} + \frac{e^{-\frac{\|\mathbf{m}\|^2}{2\sigma^2}}}{2\pi j} \sum_{n=0}^{\infty} \frac{1}{p!} \left(\frac{\|\mathbf{m}\|^2}{2\sigma^2}\right)^p I_{n+p}$$
 (15)

From (13) and (14), we now have

$$I_n = \int_{-\infty}^{\infty} \frac{1}{(1 - 2jt)^{\frac{1}{2}}} \left[\frac{1}{t} - 2j\sigma^2 \sum_{k=1}^n \frac{1}{(1 + 2j\sigma^2 t)^k} \right] dt, \quad (16)$$

which, after changing the order of the integral and the summation can be written as

$$I_n = \int_{-\infty}^{\infty} \frac{dt}{t(1 - 2jt)^{\frac{1}{2}}} - 2j\sigma^2 \sum_{k=1}^{n} \int_{-\infty}^{\infty} \frac{dt}{(1 - 2jt)^{\frac{1}{2}}(1 + 2j\sigma^2t)^k}.$$

In the above, letting

$$J = \int_{-\infty}^{\infty} \frac{dt}{t(1 - 2it)^{\frac{1}{2}}},$$
 (17)

$$J_k = 2j\sigma^2 \int_{-\infty}^{\infty} \frac{dt}{(1-2jt)^{\frac{1}{2}}(1+2j\sigma^2t)^k}, \quad (18)$$

we can write (14) as

$$I_n = J - \sum_{k=1}^{n} J_k. (19)$$

In the next section, we first show that J and J_k can be reduced to simple real and contour integrals respectively and then solve them.

III. SOLVING FOR I_n

A. The Real Integral

In (17), through a change of variables (from t to -t), we obtain

$$J = -\int_{-\infty}^{\infty} \frac{dt}{t(1+2jt)^{\frac{1}{2}}}.$$
 (20)

Adding the expressions for J in (17) and (20),

$$2J = \int_{-\infty}^{\infty} \frac{dt}{t(1-2jt)^{\frac{1}{2}}} - \int_{-\infty}^{\infty} \frac{dt}{t(1+2jt)^{\frac{1}{2}}}$$
(21)
$$= \int_{-\infty}^{\infty} \frac{1}{t} \left[\frac{(1+2jt)^{\frac{1}{2}} - (1-2jt)^{\frac{1}{2}}}{(1+4t^2)^{\frac{1}{2}}} \right] dt.$$
(22)

Multiplying the numerator and denominator of the integrand in (22) by $\left[(1+2jt)^{\frac{1}{2}} - (1-2jt)^{\frac{1}{2}} \right]$,

$$J = \frac{1}{2} \int_{-\infty}^{\infty} \frac{1}{t} \left[\frac{(1+2jt) - (1-2jt)}{(1+4t^2)^{\frac{1}{2}} \left\{ (1+2jt)^{\frac{1}{2}} + (1-2jt)^{\frac{1}{2}} \right\}} \right] dt. \quad (23)$$

Cancelling out all common factors.

$$J = 2j \int_{-\infty}^{\infty} \frac{dt}{(1+4t^2)^{\frac{1}{2}} \left\{ (1+2jt)^{\frac{1}{2}} + (1-2jt)^{\frac{1}{2}} \right\}}.$$
 (24)

In the above equation, we note that $(1-2j)^{\frac{1}{2}}$ is the complex cojugate of $(1+2j)^{\frac{1}{2}}$. Hence, the integrand in (24) is real as well as an even function of t. Thus, we get (see Appendix)

$$J = 4j \int_0^\infty \frac{dt}{(1+4t^2)^{\frac{1}{2}} \left\{ (1+2jt)^{\frac{1}{2}} + (1-2jt)^{\frac{1}{2}} \right\}}$$
(25)
$$= j\pi.$$
(26)

B. The contour integral

The integral in (18) can be solved easily if it can be converted to a contour integral. Towards this end, we state the following Lemma [8].

Lemma 3.1: Let g(x) be a function of a real variable x such that |g(x)| has a denominator different from zero for all real x and is of degree in excess of a unit higher than the degree of the numerator. Then

$$\int_{-\infty}^{\infty} g(x)dx = \int_{C} g(z)dz,$$
(27)

where C is a semicircle in the complex upper half-plane whose diameter is the real-axis and the integration is in the anti-clockwise sense.

For the integrand in (18), $k \ge 1$ and the degree of the denominator is greater than that of the numerator by $k + \frac{1}{2}$. From Lemma 3.1, we get

$$J_{k} = 2j\sigma^{2} \int_{C} \frac{dz}{(1 - 2jz)^{\frac{1}{2}}(1 + 2j\sigma^{2}z)^{k}}$$

$$= (2j\sigma^{2})^{1-k} \int_{C} \frac{dz}{(1 - 2jz)^{\frac{1}{2}}(z - \frac{j}{2-2})^{k}}.$$
(28)

We now present a formula for finding the derivatives of an analytic function [8] and subsequently use it to evaluate I_k .

Lemma 3.2: If g(z) is analytic in a domain D, then it has derivatives of all orders in D which are then also analytic functions in D. The value of the (k-1)th derivative at a point z_0 in D is given by the formula

$$g^{(k-1)}(z_0) = \frac{(k-1)!}{2\pi j} \int_L \frac{g(z)}{(z-z_0)^k} \quad (k=1,2,\ldots); \quad (29)$$

where L is any simple closed path in D which encloses z_0 an whose full interior belongs to D; the curve is traversed in the counterclockwise sense and $g^{(0)}(z_0) = g(z_0)$, by definition. The function

$$f(z) = \frac{1}{(1 - 2iz)^{\frac{1}{2}}} \tag{30}$$

is analytic in the upper half-plane and C is a closed path in it. Since (28) can be written as

$$J_k = (2j\sigma^2)^{1-k} \int_C \frac{f(z)}{(z - \frac{j}{2\sigma^2})^k} dz,$$
 (31)

and the point $\frac{j}{2\sigma^2}$ lies within C, using Lemma 3.2, we obtain

$$J_k = \frac{2\pi j (2j\sigma^2)^{1-k}}{(k-1)!} \frac{d^{k-1}}{dz^{k-1}} \left[\frac{1}{(1-2jz)^{\frac{1}{2}}} \right]_{z=\frac{j}{2\sigma^2}}$$
(32)

which, after some simplification, yields

$$J_{k} = \frac{2\pi j\sigma}{\sqrt{1+\sigma^{2}}}, k = 1,$$

$$= \frac{2\pi j\sigma}{(k-1)!} \frac{\frac{1}{2} \frac{3}{2} \dots \frac{(2k-3)}{2}}{(1+\sigma^{2})^{k-\frac{1}{2}}}, \quad 1 < k \le n.$$
(33)

Substituting the expressions obtained in (26) and (33) in (19), we get

$$I_{n} = j\pi - \frac{2\pi j\sigma}{\sqrt{1+\sigma^{2}}} \left[1 + \sum_{k=2}^{n} \frac{\frac{1}{2} \cdot \frac{3}{2} \cdots \frac{(2k-3)}{2}}{(k-1)!(1+\sigma^{2})^{k-1}} \right]$$
(34)
$$= j\pi - \frac{2\pi j\sigma}{\sqrt{1+\sigma^{2}}} \left[1 + \sum_{k=1}^{n-1} \frac{\frac{1}{2} \cdot \frac{3}{2} \cdots \frac{(2k-1)}{2}}{k!(1+\sigma^{2})^{k}} \right]$$

$$= 2\pi j \left[\frac{1}{2} - \frac{\sigma}{\sqrt{1+\sigma^{2}}} \sum_{k=0}^{n-1} {2k \choose k} \left\{ \frac{1}{4(1+\sigma^{2})} \right\}^{k} \right].$$

IV. CLOSED FORM EXPRESSION FOR $E[Q(\|\mathbf{X}\|)]$ From (15) and (34),

$$P(\Delta < 0) = 1 - \frac{\sigma e^{-\frac{\|\mathbf{m}\|^2}{2\sigma^2}}}{\sqrt{1+\sigma^2}} \sum_{p=0}^{\infty} \sum_{k=0}^{n+p-1} \frac{1}{p!} \left(\frac{\|\mathbf{m}\|^2}{2\sigma^2}\right)^p \times \left(\frac{2k}{k}\right) \left\{\frac{1}{4(1+\sigma^2)}\right\}^k.$$
(35)

Let $\alpha = \frac{\|\mathbf{m}\|}{2\sigma^2}$ and $\beta = \frac{1}{1+\sigma^2}$. Then, changing the indices of summation,

$$P(\Delta < 0) = 1 - e^{-\alpha} \sqrt{1 - \beta} (A + B)$$
 (36)

where1

$$A = \sum_{p=0}^{\infty} \sum_{k=0}^{p} {2k \choose k} \left\{ \frac{\beta}{4} \right\}^k \frac{\alpha^p}{p!}, \tag{37}$$

$$B = \sum_{p=0}^{\infty} \sum_{k=p}^{n+p-1} {2k \choose k} \left\{ \frac{\beta}{4} \right\}^k \frac{\alpha^p}{p!}.$$
 (38)

We define the factorial function [9] as

$$(\gamma)_q = \prod_{r=1}^q (\gamma + r - 1), \quad (\gamma)_0 = 1, \gamma \neq 0,$$
 (39)

where q is a positive integer.

A. The B series

Since

$$\binom{2k}{k} = \frac{4^k \left(\frac{1}{2}\right)_k}{k!}, \tag{40}$$

we obtain

$$B = \sum_{p=0}^{\infty} \sum_{k=p}^{n+p-1} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} \frac{\alpha^p}{p!}.$$
 (41)

Changing the limits of summation in (41),

$$B = \sum_{p=0}^{\infty} \sum_{k=0}^{n-1} \frac{\left(\frac{1}{2}\right)_{k+p} \beta^{k+p}}{(k+p)!} \frac{\alpha^{p}}{p!}$$

$$= \sum_{k=0}^{n-1} \frac{\left(\frac{1}{2}\right)_{k} \beta^{k}}{k!} \sum_{p=0}^{\infty} \frac{\left(\frac{1}{2} + k\right)_{p}}{(k+1)_{p}} \frac{(\alpha\beta)^{p}}{p!}$$

$$= \sum_{k=0}^{n-1} \binom{2k}{k} \left(\frac{\beta}{4}\right)^{k} {}_{1}F_{1}\left(\frac{1}{2} + k; k+1; \alpha\beta\right),$$
(42)

where ${}_{1}F_{1}(a;b;x)$ is the confluent hypergeometric function [9]. According to Kummer's formula for the confluent hypergeometric function,

$$_{1}F_{1}(a;b;x) = e^{x} {}_{1}F_{1}(b-a;b;-x).$$
 (43)

Using this result in (42), we obtain

$$B = \exp(\alpha \beta) \sum_{k=0}^{n-1} {2k \choose k} \left(\frac{\beta}{4}\right)^k {}_1F_1\left(\frac{1}{2}; k+1; -\alpha \beta\right). \tag{44}$$

¹We assume that all the infinite series considered henceforth converge.

B. The A series

We rewrite (37) as

$$A = \sum_{p=0}^{\infty} \sum_{k=0}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} \frac{\alpha^p}{p!} - \sum_{p=0}^{\infty} \sum_{k=p+1}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} \frac{\alpha^p}{p!}$$
(45)

In the above.

$$\sum_{p=0}^{\infty} \sum_{k=0}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} \frac{\alpha^p}{p!} = \left(\sum_{p=0}^{\infty} \frac{\alpha^p}{p!}\right) \left(\sum_{k=0}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!}\right). \tag{46}$$

Since $|\beta| < 1$, the second sum on the right hand side of (46) is the binomial series, i.e.,

$$\sum_{k=0}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} = (1-\beta)^{-\frac{1}{2}}.$$
 (47)

Thus,

$$A = \frac{e^{\alpha}}{\sqrt{1-\beta}} - S,\tag{48}$$

where

$$S = \sum_{p=0}^{\infty} \sum_{k=p+1}^{\infty} \frac{\left(\frac{1}{2}\right)_k \beta^k}{k!} \frac{\alpha^p}{p!}$$

$$= \sum_{p=0}^{\infty} \sum_{k=1}^{\infty} \frac{\left(\frac{1}{2}\right)_{k+p} \beta^{k+p}}{(k+p)!} \frac{\alpha^p}{p!}.$$
(49)

Following the steps in (42), (49) can be written as

$$S = \sum_{k=1}^{\infty} {2k \choose k} \left(\frac{\beta}{4}\right)^k {}_1F_1\left(\frac{1}{2} + k; k+1; \alpha\beta\right). \tag{50}$$

The above infinite series has a closed form expression [3]

$$S = \frac{2\exp(\frac{\alpha\beta}{2})}{\sqrt{1-\beta}} \tag{51}$$

$$\times \left[\exp\left\{ \frac{\alpha}{2} (1+\beta) \right\} Q_1(u,w) - \frac{1}{2} (1+\sqrt{1-\beta}) I_0\left(\frac{\alpha\beta}{2}\right) \right], \quad (52)$$

where

$$u = \sqrt{\frac{\alpha}{2} (2 - \beta) - \frac{2}{\beta} \sqrt{1 - \beta}}$$

$$w = \sqrt{\frac{\alpha}{2} (2 - \beta) + \frac{2}{\beta} \sqrt{1 - \beta}},$$
(53)

and $Q_1(u,w)$ is the Marcum Q-function [5]. Substituting (51) in (48) gives us a closed form expression for A. Since we already have a compact expression for B in (44), replacing the infinite series for A and B in (35) by their respective closed form expressions, and noting from (2) and (3) that

$$E[Q(\|\mathbf{X}\|)] = \frac{1}{2}P(\Delta < 0),$$
 (54)

we obtain an exact expression for $E[Q(||\mathbf{X}|)]^2$.

Corollary: When X is zero mean, from (34) and (35), we obtain

$$P(\Delta < 0) = \frac{1}{2} + \frac{I_n}{2\pi j}$$

$$= 1 - \frac{\sigma}{\sqrt{1 + \sigma^2}} \sum_{k=0}^{n-1} {2k \choose k} \left\{ \frac{1}{4(1 + \sigma^2)} \right\}^k.$$
(55)

Substituting the above in (54) leads to the well known result [2]

$$E[Q(\|\mathbf{X}\|)] = \frac{1}{2} \left[1 - \frac{\sigma}{\sqrt{1 + \sigma^2}} \sum_{k=0}^{n-1} {2k \choose k} \left\{ \frac{1}{4(1 + \sigma^2)} \right\}^k \right]$$
(56)

V. Example: Average Probability of Error for Nakagami-m fading channels

If a random variable α is Nakagami-m distributed, the random variable $\gamma = \frac{\alpha^2 \varepsilon_b}{N_0}$ has the probability density function [5]

$$p(\gamma) = \frac{m^m}{\Gamma(m)\bar{\gamma}} \gamma^{m-1} e^{-m\gamma/\bar{\gamma}}, \tag{57}$$

where $\bar{\gamma} = \frac{E(\alpha^2)\varepsilon_b}{N_0}$. For fading channels, the average probability of error is given by

$$P_e = \int_0^\infty Q(a\sqrt{\gamma})p_\gamma(\gamma)d\gamma,\tag{58}$$

where a is a constant that depends on the specific modulation/detection combination [1]. We note that γ has the same distribution as R (when \mathbf{X} has zero mean) with $\sigma^2 = \frac{\bar{\gamma}}{2m}$ when m is an integer [5]. Thus, after accounting for the constant a in (58), the average probability of error for a Nakagami-m fading channel is obtained as

$$P_e = \frac{1}{2} \left[1 - \sqrt{\frac{a^2 \bar{\gamma}}{2m + a^2 \bar{\gamma}}} \sum_{k=0}^{m-1} {2k \choose k} \left\{ \frac{2m}{4(2m + a^2 \bar{\gamma})} \right\}^k \right]$$
(59)

by substituting $\sigma^2 = \frac{a^2 \bar{\gamma}}{2m}$ and replacing n by m in (56). We note that exactly the same result has been arrived at using a different approach in [1], equation (5.18).

APPENDIX

Let $1+2jt=re^{j\theta}$, where $r=(1+4t^2)^{\frac{1}{2}}$ and $\cos\theta=\frac{1}{r}$. Since

$$(1+2jt)^{\frac{1}{2}} + (1-2jt)^{\frac{1}{2}} = 2r^{\frac{1}{2}}\cos\frac{\theta}{2},\tag{60}$$

and

$$\cos\frac{\theta}{2} = \sqrt{\frac{1}{2}(1+\cos\theta)}$$

$$= \sqrt{\frac{1}{2}\left(1+\frac{1}{r}\right)},$$
(61)

the integrand in (25)

$$\frac{1}{(1+4t^2)^{\frac{1}{2}}\left\{(1+2jt)^{\frac{1}{2}}+(1-2jt)^{\frac{1}{2}}\right\}} = \frac{1}{2r\sqrt{\frac{1+r}{2}}}.$$
 (62)

²Lindsey, in [3], using a fundamentally different approach, has obtained a closed form expression in a completely different form.

Hence,

$$J = 4j \int_0^\infty \frac{dt}{(1+4t^2)^{\frac{1}{2}} \left\{ (1+2jt)^{\frac{1}{2}} + (1-2jt)^{\frac{1}{2}} \right\}}$$
(63)
$$= 2\sqrt{2}j \int_0^\infty \frac{dt}{(1+4t^2)^{\frac{1}{2}} \left\{ 1 + (1+4t^2)^{\frac{1}{2}} \right\}^{\frac{1}{2}}}.$$

Substituting $t = \frac{\tan \phi}{2}$ in the above,

$$J = \sqrt{2}j \int_0^{\frac{\pi}{2}} \frac{\sec \phi}{\sqrt{1 + \sec \phi}} d\phi. \tag{64}$$

Let $\sqrt{1 + \sec \phi} = \sqrt{2} \sec \psi$. Then

$$\frac{\sec\phi}{\sqrt{1+\sec\phi}}d\phi = \sqrt{2}d\psi. \tag{65}$$

The integral in (64) now becomes

$$J = 2j \int_0^{\frac{\pi}{2}} d\psi$$
 (66)
= $j\pi$.

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